EXTREMAL DICHOTOMY FOR UNIFORMLY HYPERBOLIC SYSTEMS

MARIA CARVALHO, ANA CRISTINA MOREIRA FREITAS, JORGE MILHAZES FREITAS, MARK HOLLAND, AND MATTHEW NICOL

ABSTRACT. We consider the extreme value theory of a hyperbolic toral automorphism $T:\mathbb{T}^2\to\mathbb{T}^2$ showing that if a Hölder observation ϕ which is a function of a Euclidean-type distance to a non-periodic point ζ is strictly maximized at ζ then the corresponding time series $\{\phi\circ T^i\}$ exhibits extreme value statistics corresponding to an iid sequence of random variables with the same distribution function as ϕ and with extremal index one. If however ϕ is strictly maximized at a periodic point q then the corresponding time-series exhibits extreme value statistics corresponding to an iid sequence of random variables with the same distribution function as ϕ but with extremal index not equal to one. We give a formula for the extremal index (which depends upon the metric used and the period of q). These results imply that return times are Poisson to small balls centered at non-periodic points and compound Poisson for small balls centered at periodic points.

1. Introduction

Suppose we have a time-series $\{X_i\}$ of real-valued random variables defined on a probability space (Ω, P) . Extreme value theory concerns the statistical behavior of the derived time series of maxima $M_n := \max\{X_1, \ldots, X_n\}$. There is a well-developed theory for iid random variables [24, 26, 14] and these references also consider the case of dependent random variables under certain assumptions. In the setting of ergodic maps $T: X \to X$ of a probability space (X, μ) modeling deterministic physical phenomena we are concerned with time-series of the form $\{\phi \circ T^i\}$ where $\phi: X \to \mathbb{R}$ is an observation of some regularity, usually Hölder. In the literature, as in this paper, it is customary to assume that ϕ is a function of the distance d(x, p) to a distinguished point ζ for some metric d so that $\phi(x) = f(d(x, \zeta))$ for $x \in X$. For convenience the observation $\phi(x) = -\log d(x, \zeta)$ is often

Date: January 27, 2015.

²⁰⁰⁰ Mathematics Subject Classification. 37A50, 60G70, 37B20, 60G10, 37C25.

Key words and phrases. Extreme Value Theory, Return Time Statistics, Stationary Stochastic Processes, Metastability.

ACMF was partially supported by FCT grant SFRH/BPD/66174/2009. JMF was partially supported by FCT grant SFRH/BPD/66040/2009. Both these grants are financially supported by the program POPH/FSE. ACMF and JMF are supported by FCT (Portugal) project PTDC/MAT/120346/2010, which is financed by national and European structural funds through the programs FEDER and COMPETE. MC, ACMF and JMF are also supported by CMUP, which is financed by FCT (Portugal) through the programs POCTI and POSI, with national and European structural funds, under the project PEst-C/MAT/UI0144/2013. JMF would like to thank Mike Todd for helpful comments and suggestions.

used, but scaling relations translate extreme value results for one functional to another quite easily [9, 19]. Since the function $\phi \circ T^i(x)$ detects the closer $T^i(x)$ is to ζ , there is a natural relation between extreme value statistics for the time series $\{\phi \circ T^i\}$ and return time statistics to nested balls about ζ [4, 10]. In fact, if (X, d) is a Riemannian manifold and μ is absolutely continuous with respect to the volume measure then return time statistics may be deduced from extreme value statistics and vice-versa, so in some sense they are equivalent though proved with different techniques.

Due to the close relation with return time statistics one expects a difference in the extreme value theory of observations centered at periodic points to those centered at generic points. We refer to the early paper of Hirata [18]. What is perhaps surprising is that in many cases of uniformly expanding smooth one dimensional systems there is a strict dichotomy between the behavior of observations maximized at periodic points and other non-periodic points. Suppose ϕ is maximized at a point $\zeta \in X$ and define a sequence of constants u_n by the requirement that $\lim_{n\to\infty} n\mu(\phi>u_n)=\tau$, where $\tau\in\mathbb{R}$. It has been shown [7, 10, 22] that for a variety of smooth one-dimensional chaotic systems this implies that $\lim_{n\to\infty}\mu(M_n\leq$ u_n) = $e^{-\vartheta \tau}$ where $0 < \vartheta \le 1$ is called the extremal index and roughly measures the clustering of exceedances of the maxima (more precisely ϑ^{-1} is the expected number of observed exceedances if one exceedance is observed). It turns out that $\vartheta = 1$ if ζ is not periodic and $\vartheta < 1$ otherwise (a precise expression for ϑ may often be given as a function of the multiplier derivative of T along the periodic orbit). For higher dimensional smooth chaotic systems a similar dichotomy is expected (the presence of discontinuities complicates matters [1]) but so far has not been established. In this paper we establish such a dichotomy for hyperbolic toral automorphisms, give an expression for the extremal index and find that it also depends upon the metric used. This has immediate implications for the return time statistics. We have Poissonian return time statistics to non-periodic points and a compound Poissonian distribution to periodic points (the form of the compound distribution is related to the extremal index, as well). In fact, our results are given in terms of convergence of point processes of rare events to the standard Poisson process, in the non-periodic case, and to a compound Poisson process, in the periodic case, where the multiplicity distribution is also seen to depend on the metric used. We remark that the convergence of the point processes is actually stronger than establishing a Poissonian distributional limit for the number of visits to shrinking target sets.

Earlier results for invertible dynamical systems include work by Denker, Gordin and Sharova, [5], who proved that, for toral automorphisms, the number of visits to shrinking neighbourhoods of non-periodic points converge (when properly normalised) to a Poisson distribution and work by Dolgopyat [6], who also proved Poissonian return time statistics for non-periodic points in uniformly partially hyperbolic dynamical systems and noted possible generalizations to the case of periodic points. There have been a variety of recent results on extreme value and return time statistics for invertible dynamical systems, both uniformly and non-uniformly hyperbolic, but restricted to observations maximized at generic points [16, 19, 3, 13, 27, 15].

2. The setting

Let $L: \mathbb{R}^2 \to \mathbb{R}^2$ be a 2×2 matrix with integer entries, with determinant equal to 1 and without eigenvalues of absolute value 1. The eigenvalues of L are irrational numbers λ and $1/\lambda$ satisfying $|\lambda| > 1$; the corresponding eigenspaces, E^s and E^u , are lines with irrational slope and the family of lines parallel to these subspaces is invariant under L. Moreover, L contracts vectors in E^s and expands in E^u . As $L(\mathbb{Z}^2) = \mathbb{Z}^2$, this linear map induces an invertible smooth transformation T on the flat torus $\mathbb{T}^2 = \mathbb{R}^2/\mathbb{Z}^2$ through the canonical projection. The subspaces E^s and E^u project on dense curves which intersect densely and are T invariant; that is, the torus is the homoclinic class of the (unique) fixed point of T and a global attractor.

After finding a norm in \mathbb{R}^2 that makes L a uniform contraction in E^s and a uniform expansion in E^u , we project the Riemannian metric generated by this norm to \mathbb{T}^2 and consider the invariant splitting at each point into one-dimensional affine spaces parallel to the subspaces E^s and E^u . This way, T becomes an Anosov diffeomorphism: the hyperbolic splitting at each point of \mathbb{T}^2 is obtained by translating to that point the eigenspaces of the matrix. Through each point $x \in \mathbb{T}^2$ there is a local stable manifold $W^s_{\epsilon}(x)$ of length ϵ and similarly a local unstable manifold $W^u_{\epsilon}(x)$. Besides, T is topologically mixing and its periodic points (all saddles) are dense, even though for each $q \in \mathbb{N}$ the set of periodic points with period q is finite.

As the determinant of L is equal to 1, the Riemannian structure induces a Lebesgue measure m on \mathbb{T}^2 which is invariant by T. The probability measure m is Bernoulli and it is also the unique equilibrium state of the Hölder map $x \in \mathbb{T}^2 \to \varphi^u(x) = -\log |\det DT_{E^u(x)}|$, whose pressure is zero. This means that the metric entropy of m is given by $h_m(T) = -\int \varphi^u dm = \log(\lambda)$.

Unless said otherwise, we assume that we are using the Euclidean metric in \mathbb{R}^2 , which we project to \mathbb{T}^2 and denote by $d(\cdot,\cdot)$. More precisely, if (x,y) denote the local coordinates of the point $z \in \mathbb{T}^2$ then

$$d(z,0) = \sqrt{x^2 + y^2}. (2.1)$$

Given a non-empty set A and a point x, we denote by \bar{A} the closure of the set A and define the distance of x to A by $\operatorname{dist}(x,A) := \inf\{d(x,y): y \in A\}$.

Remark 2.1. Let A be a set homeomorphic to a ball whose boundary is a piecewise differentiable Jordan curve. For $\varepsilon > 0$, let $D_{\varepsilon} := \{x \in \mathbb{T}^2 : \operatorname{dist}(x, \bar{A}) \leq \varepsilon\}$. Then $m(D_{\varepsilon} \setminus A) = \varepsilon |\partial A| + o(\varepsilon)$, where $|\cdot|$ denotes the length of the curve ∂A .

Recall that a continuous function $\varphi: \mathbb{T}^2 \to \mathbb{R}$ is α -Hölder if there are constants C > 0 and $\alpha > 0$ such that $|\varphi(x) - \varphi(y)| \le C d(x,y)^{\alpha}$. For $\alpha \in]0,1]$, the space of α -Hölder continuous functions with the norm

$$\|\varphi\|_{\alpha} = \|\varphi\|_{C^0} + \sup_{x \neq y} \left\{ \frac{|\varphi(x) - \varphi(y)|}{d(x, y)^{\alpha}} \right\}$$

is a Banach space. And, given $k \in \mathbb{N}$ and $\alpha \in]0,1[$, any α -Hölder continuous function φ on the torus admits a Lipschitz approximation φ_k such that $\|\varphi - \varphi_k\|_{\infty} = O(k^{-\frac{\alpha}{1-\alpha}})$. From the computations in [2, Section 1.26] one can deduce that the diffeomorphism T has exponential decay of correlations with respect to α -Hölder functions, for any $\alpha \in]0,1]$. More precisely, there exist constants C>0 and $0<\theta<1$ such that, for all α -Hölder maps ϕ and ψ , we have

$$\left| \int \phi \left(\psi \circ T^n \right) dm - \int \phi dm \int \psi dm \right| \le C \theta^n \|\phi\|_{\alpha} \|\psi\|_{\alpha}. \tag{2.2}$$

And if ψ is constant on local stable leaves then

$$\left| \int \phi \left(\psi \circ T^n \right) dm - \int \phi dm \int \psi dm \right| \le C \theta^n \|\phi\|_{\alpha} \|\psi\|_{\infty}. \tag{2.3}$$

Among C^2 Anosov diffeomorphisms the subset of the ones that admit no invariant measure absolutely continuous with respect to m is open and dense. However, it is known that any Anosov diffeomorphism of the torus is conjugate to a linear model; see [25]. The conjugacy H (and its inverse) is Hölder continuous [21] and carries over the measure m to a probability measure $\mu = H_*(m)$ on \mathbb{T}^2 which is invariant and mixing by the original Anosov diffeomorphism. In particular, the exponential decay of correlations (2.2) is still valid for μ , although corresponding to different values of C, θ and α . We remark that the conditions we will analyze in the next section, to ensure the existence of an Extreme Value Law, are also invariant under conjugacy.

The starting point of our analysis is a stationary stochastic process X_0, X_1, X_2, \ldots arising from the system described above in the following way. We fix some point $\zeta \in \mathbb{T}^2$ and let $\varphi : \mathbb{T}^2 \to \mathbb{R} \cup \{+\infty\}$ be given by

$$\varphi(z) = -\log(d(z,\zeta)),\tag{2.4}$$

where d denotes the projection of the usual Euclidean metric in \mathbb{R}^2 to \mathbb{T}^2 , as in (2.1). We define $X_0 = \varphi$ and for all $i \in \mathbb{N}$

$$X_i = \varphi \circ T^i. \tag{2.5}$$

Since m is an invariant probability measure on \mathbb{T}^2 , then the stochastic process X_0, X_1, X_2, \ldots just defined is stationary.

Note that $\{X_0 > u\} = B_{e^{-u}}(\zeta)$, where $B_{e^{-u}}(\zeta)$ denotes the ball centered at ζ of radius e^{-u} , in the Euclidean metric. In other words, having an exceedance of an high threshold u, at time j, means that, at the jth iterate, the orbit enters the ball around ζ of radius e^{-u} .

We remark that we could use a function different from $-\log$ in the definition of φ . In fact, any not too irregular function, achieving a global maximum at 0 and invertible in a vicinity of 0, would work. See [9, Section 1.1] for conditions on this function. For definiteness, here we will use $-\log$ which puts us on the Type 1 or Gumbel domain of attraction for the maxima. Our goal is to show the existence of an EVL and for that purpose the analysis does not change much from one example to another.

In order to obtain a non-degenerate limit for the distribution of $M_n := \max\{X_0, \dots, X_{n-1}\}$, one usually uses normalizing sequences $(u_n)_{n\in\mathbb{N}}$ such that the average number of exceedances among the first n observations is some constant $\tau \geq 0$, when n is large enough, i.e., the sequence $(u_n)_{n\in\mathbb{N}}$ satisfies:

$$n \, m(X_0 > u_n) \xrightarrow[n \to \infty]{} \tau \ge 0. \tag{2.6}$$

In the case here, if we take $u_n = \frac{1}{2} \log(\pi n) - \frac{1}{2} \log \tau$, then conditions (2.6) is satisfied.

We are now in condition of stating our first result which asserts a dichotomy regarding the possible limit distributions for M_n depending on the periodicity or aperiodicity of the point ζ .

Theorem 2.2. Let $T: \mathbb{T}^2 \to \mathbb{T}^2$ be an Anosov linear diffeomorphism. Fix $\zeta \in \mathbb{T}^2$ and define the stochastic process X_0, X_1, X_2, \ldots as in (2.5). Let $(u_n)_{n \in \mathbb{N}}$ be a sequence satisfying (2.6) for some $\tau \geq 0$. Then

- (1) if ζ is not periodic, we have $\lim_{n\to\infty} m(M_n \leq u_n) = e^{-\tau}$.
- (2) if ζ is periodic of prime period q, we have $\lim_{n\to\infty} m(M_n \leq u_n) = e^{-\vartheta \tau}$, where $\vartheta = \frac{2}{\pi} \left(\arcsin \frac{|\lambda|^q}{\sqrt{|\lambda|^{2q}+1}} \arcsin \frac{1}{\sqrt{|\lambda|^{2q}+1}} \right)$.

In contrast to the examples of uniformly expanding systems (with no contracting directions), the metric used plays an important role in the computation of the EI. To illustrate this, let e^u , e^s denote unit eigenvectors associated to the eigenvalues λ and $1/\lambda$, respectively; thus $\{e^u, e^s\}$ forms a normal basis of \mathbb{R}^2 . Let $z = (x^u, x^s)$, where (x^u, x^s) stands for the coordinates of z in this basis. Define the metric d^* by

$$d^*(z,0) = \max\{|x^u|, |x^s|\}$$
(2.7)

and then project it to \mathbb{T}^2 .

Theorem 2.3. Let $T: \mathbb{T}^2 \to \mathbb{T}^2$ be an Anosov linear diffeomorphism. Fix $\zeta \in \mathbb{T}^2$ and define the stochastic process X_0, X_1, X_2, \ldots as in (2.5) for the observable φ as in (2.4) but evaluated with the metric d^* instead. Let $(u_n)_{n \in \mathbb{N}}$ be a sequence satisfying (2.6) for some $\tau \geq 0$. Then

- (1) if ζ is not periodic, we have $\lim_{n\to\infty} m(M_n \leq u_n) = e^{-\tau}$.
- (2) if ζ is periodic of prime period q, we have $\lim_{n\to\infty} m(M_n \leq u_n) = e^{-\vartheta \tau}$, where $\vartheta = 1 |\lambda|^{-q}$.

3. The approach and dependence conditions

In order to prove the existence of EVLs in a dynamical systems context, there are a couple of conditions on the dependence structure of the stochastic process that if verified allow us to obtain such distributional limits. These conditions are motivated by the conditions $D(u_n)$ and $D'(u_n)$ of Leadbetter but were adapted to the dynamical setting and further

developed both in the absence of clustering, such as in [4, 8, 19], and in the presence of clustering in [10]. Very recently, in [12], the authors provided some more general conditions, called $\mathcal{A}(u_n)$ and $\mathcal{A}'_q(u_n)$, which subsumed the previous ones and allowed them to address both the presence $(q \geq 1)$ and the absence (q = 0) of clustering. To distinguish these conditions the authors used a Cyrillic D to denote them. We recall these conditions here.

Given a sequence $(u_n)_{n\in\mathbb{N}}$ of real numbers satisfying (2.6) and $q\in\mathbb{N}_0$, set

$$A_n^{(q)} := \{X_0 > u_n, X_1 \le u_n, \dots, X_q \le u_n\}.$$

For $s, \ell \in \mathbb{N}$ and an event B, let

$$\mathcal{W}_{s,\ell}(B) = \bigcap_{i=s}^{s+\ell-1} T^{-i}(B^c). \tag{3.1}$$

Condition $(\mathcal{A}_q(u_n))$. We say that $\mathcal{A}(u_n)$ holds for the sequence X_0, X_1, \ldots if, for every $\ell, t, n \in \mathbb{N}$

$$\left| \mathbb{P}\left(A_n^{(q)} \cap \mathcal{W}_{t,\ell}\left(A_n^{(q)} \right) \right) - \mathbb{P}\left(A_n^{(q)} \right) \mathbb{P}\left(\mathcal{W}_{0,\ell}\left(A_n^{(q)} \right) \right) \right| \le \gamma(q, n, t), \tag{3.2}$$

where $\gamma(q, n, t)$ is decreasing in t and there exists a sequence $(t_n)_{n \in \mathbb{N}}$ such that $t_n = o(n)$ and $n\gamma(q, n, t_n) \to 0$ when $n \to \infty$.

For some fixed $q \in \mathbb{N}_0$, consider the sequence $(t_n)_{n \in \mathbb{N}}$ given by condition $\mathcal{L}_q(u_n)$ and let $(k_n)_{n \in \mathbb{N}}$ be another sequence of integers such that

$$k_n \to \infty$$
 and $k_n t_n = o(n)$. (3.3)

Condition $(A'_q(u_n))$. We say that $A'_q(u_n)$ holds for the sequence X_0, X_1, \ldots if there exists a sequence $(k_n)_{n\in\mathbb{N}}$ satisfying (3.3) and such that

$$\lim_{n \to \infty} n \sum_{j=1}^{\lfloor n/k_n \rfloor} \mathbb{P}\left(A_n^{(q)} \cap T^{-j}\left(A_n^{(q)}\right)\right) = 0. \tag{3.4}$$

We note that, when q = 0, condition $\mathcal{A}'_q(u_n)$ corresponds to condition $\mathcal{D}'(u_n)$ from [23].

Now let

$$\vartheta = \lim_{n \to \infty} \vartheta_n = \lim_{n \to \infty} \frac{\mathbb{P}(A_n^{(q)})}{\mathbb{P}(U_n)}.$$
(3.5)

From [12, Corollary 2.4], it follows that if the stochastic process X_0, X_1, \ldots satisfies both conditions $\mathcal{A}_q(u_n)$ and $\mathcal{A}'_q(u_n)$, the limit in (3.5) exists and

$$\lim_{n\to\infty} \mathbb{P}(M_n \le u_n) = e^{-\vartheta \tau}.$$

Hence, the strategy to prove Theorems 2.2 and 2.3 is to show that conditions $\mathcal{L}_0(u_n)$ and $\mathcal{L}'_0(u_n)$ hold when ζ is a non periodic point, and that conditions $\mathcal{L}_q(u_n)$ and $\mathcal{L}'_q(u_n)$ hold when ζ is a periodic point of prime period q. In fact, we only check these conditions for the usual Euclidean metric as in the context of Theorem 2.2, which is technically harder, leaving the necessary adjustments when dealing with the adapted metric d^* for the reader.

4. Checking condition $\mathcal{A}_q(u_n)$

Let D be a set homeomorphic to a ball whose boundary is piecewise smooth, so $m(\partial D) = 0$, and define

$$H_k(D) = \{x \in D : T^k(W_1^s(x)) \cap \partial D \neq \emptyset\}.$$

Proposition 4.1. There exist constants C > 0 and $0 < \tau_1 < 1$ such that, for all k,

$$m(H_k(D)) \le C\tau_1^k. \tag{4.1}$$

Proof. As a consequence of the uniform contraction within local stable manifolds, there exists $C_1 > 0$ such that dist $(T^n(x), T^n(y)) \leq C_1/|\lambda|^n$ for all $y \in W_1^s(x)$. In particular, this implies that $|T^k(W_1^s(x))| \leq C_1/|\lambda|^k$. Therefore, for every $x \in H_k(D)$, the leaf $T^k(W_1^s(x))$ lies in an annulus of width $2/|\lambda|^k$ around ∂D . By Remark 2.1 and the invariance of m, the result follows with $\tau_1 = 1/|\lambda|$.

Lemma 4.2. Suppose $\Phi: M \to \mathbb{R}$ is a Lipschitz map and Ψ is the indicator function

$$\Psi := \mathbf{1}_{\mathscr{W}_{0,\ell}\left(A_n^{(q)}\right)}.$$

Then for all $j \geq 0$

$$\left| \int \Phi \left(\Psi \circ T^{j} \right) dm - \int \Phi dm \int \Psi dm \right| \leq \mathcal{C} \left(\|\Phi\|_{\infty} \ \tau_{1}^{\lfloor j/2 \rfloor} + \|\Phi\|_{Lip} \ \theta^{\lfloor j/2 \rfloor} \right). \tag{4.2}$$

Proof. We choose for reference the fixed point ζ of T and its local unstable manifold $\tilde{\gamma}^u := W^u_1(\zeta)$. By the hyperbolic product structure, each local stable manifold $W^s_1(x)$ intersects $\tilde{\gamma}^u$ in a unique point \hat{x} . Therefore, for every map Ψ we define the function $\overline{\Psi}(x) := \Psi(\hat{x})$. Also, for every $i \in \mathbb{N}$, let $\Psi_i = \Psi \circ T^i$. We note that $\overline{\Psi}_i$ is constant along stable manifolds and that the set of points where $\overline{\Psi}_i \neq \Psi_i$ is, by definition, the set of x for which there exist $r \in \mathbb{N}_0$ and x_1, x_2 on the same local stable manifold as $T^r(x)$ such that

$$x_1 \in \mathscr{W}_{i,\ell}\left(A_n^{(q)}\right)$$

but

$$x_2 \notin \mathscr{W}_{i,\ell}\left(A_n^{(q)}\right)$$
.

Moreover, this set is contained in $\bigcup_{k=i}^{i+\ell-1} H_k(A_n^{(q)})$. If we let $i \ge \lfloor j/2 \rfloor$ then by Proposition 4.1 we have

$$m(\{\Psi_i \neq \overline{\Psi}_i\}) \leq \sum_{k=|j/2|}^{\infty} m(H_k(A_n^{(q)})) \leq \mathcal{C} \, \tau_1^{\lfloor j/2 \rfloor}.$$

By (2.3), we also get

$$\left| \int \Phi\left(\overline{\Psi}_{\lfloor j/2 \rfloor} \circ T^{j-\lfloor j/2 \rfloor} \right) dm - \int \Phi dm \int \overline{\Psi}_{\lfloor j/2 \rfloor} dm \right| \leq \mathcal{C} \|\Phi\|_{\operatorname{Lip}} \|\bar{\Psi}\|_{\infty} \ \theta^{\lfloor j/2 \rfloor}.$$

Using the identity

we obtain

$$\left| \int \Phi \left(\Psi_{\lfloor j/2 \rfloor} \circ T^{j-\lfloor j/2 \rfloor} \right) dm - \int \Phi dm \int \Psi_{\lfloor j/2 \rfloor} dm \right|$$

$$\leq \left| \int \Phi \left(\left(\Psi_{\lfloor j/2 \rfloor} - \overline{\Psi}_{\lfloor j/2 \rfloor} \right) \circ T^{j-\lfloor j/2 \rfloor} \right) dm \right| + \mathcal{C} \|\Phi\|_{\text{Lip}} \theta^{\lfloor j/2 \rfloor}$$

$$+ \left| \int \Phi dm \int \left(\overline{\Psi}_{\lfloor j/2 \rfloor} - \Psi_{\lfloor j/2 \rfloor} \right) \right| dm$$

$$\leq \mathcal{C} \left(2 \|\Phi\|_{\infty} m \left\{ \overline{\Psi}_{\lfloor j/2 \rfloor} \neq \Psi_{\lfloor j/2 \rfloor} \right\} + \|\Phi\|_{\text{Lip}} \theta^{\lfloor j/2 \rfloor} \right)$$

$$\leq \mathcal{C} \left(\|\Phi\|_{\infty} \tau_{1}^{\lfloor j/2 \rfloor} + \|\Phi\|_{\text{Lip}} \theta^{\lfloor j/2 \rfloor} \right). \tag{4.3}$$

We complete the proof by observing that $\int \Psi dm = \int \Psi_{\lfloor j/2 \rfloor} dm$, due to the *T*-invariance of *m* and the fact that $\Psi_{\lfloor j/2 \rfloor} \circ T^{j-\lfloor j/2 \rfloor} = \Psi_j = \Psi \circ T^j$.

To prove condition $\mathcal{A}(u_n)$, we will approximate the characteristic function of the set $A_n^{(q)}$ by a suitable Lipschitz function. However, this Lipschitz function will decrease sharply to zero near the boundary of the set $A_n^{(q)}$. As the estimate in Lemma 4.2 involves the Lipschitz norm, we need to bound its increase as we approach $\partial A_n^{(q)}$.

Let $A_n = A_n^{(q)}$ and $D_n := \left\{ x \in A_n^{(q)} : \operatorname{dist}\left(x, \overline{A_n^c}\right) \ge n^{-2} \right\}$, where \overline{A}_n^c denotes the closure of the complement of the set A_n . Define $\Phi_n : \mathcal{X} \to \mathbb{R}$ as

$$\Phi_n(x) = \begin{cases}
0 & \text{if } x \notin A_n \\
\frac{\operatorname{dist}(x, A_n^c)}{\operatorname{dist}(x, A_n^c) + \operatorname{dist}(x, D_n)} & \text{if } x \in A_n \setminus D_n \\
1 & \text{if } x \in D_n
\end{cases}$$
(4.4)

Note that Φ_n is Lipschitz continuous with Lipschitz constant given by n^2 . Moreover, as observed in Remark 2.1, we have $\|\Phi_n - \mathbf{1}_{A_n}\|_{L^1(m)} \leq C/n^2$ for some constant C.

It follows that

$$\left| \int \mathbf{1}_{A_{n}^{(q)}} \left(\Psi_{\lfloor j/2 \rfloor} \circ T^{j-\lfloor j/2 \rfloor} \right) dm - m(A_{n}^{(q)}) \int \Psi dm \right|$$

$$\leq \left| \int \left(\mathbf{1}_{A_{n}^{(q)}} - \Phi_{n} \right) \Psi_{\lfloor j/2 \rfloor} dm \right| + \mathcal{C} \left(\|\Phi_{n}\|_{\infty} \ j^{2} \ \tau_{1}^{\lfloor j/4 \rfloor} + \|\Phi_{n}\|_{\text{Lip}} \ \theta^{\lfloor j/2 \rfloor} \right)$$

$$+ \left| \int \left(\mathbf{1}_{A_{n}^{(q)}} - \Phi_{n} \right) dm \int \Psi_{\lfloor j/2 \rfloor} dm \right|,$$

$$(4.5)$$

and consequently

$$\left| m \left(A_n^{(q)} \cap \mathscr{W}_{j,\ell}(A_n^{(q)}) \right) - m(A_n^{(q)}) \, m \left(\mathscr{W}_{0,\ell}(A_n^{(q)}) \right) \right| \le \gamma(n,j)$$

where

$$\gamma(n,j) = \mathcal{C} \left(n^{-2} + n^2 \, \theta_1^{\lfloor j/2 \rfloor} \right)$$

and

$$\theta_1 = \max \{\tau_1, \theta\}.$$

Thus if, for instance, $j = t_n = (\log n)^5$, then $n\gamma(n, t_n) \to 0$ as $n \to \infty$. Note that we have considerable freedom of choice of j in order to ensure that the previous limit is zero; taking into account the possible applications, we chose $t_n = (\log n)^5$.

5. Checking condition $A'_q(u_n)$

Before checking the condition $\mathcal{A}'_q(u_n)$, we observe that we only need to consider the sum in (3.4) up to $t_n = (\log n)^5$ since, by the exponential decay of correlations, the sum of the remaining terms goes to 0. More precisely, we use decay of correlations to show that

$$\lim_{n \to \infty} n \sum_{j=\log^5 n}^{\lfloor n/k_n \rfloor} m(A_n^{(q)} \cap T^{-j} A_n^{(q)}) = 0.$$
 (5.1)

Let Φ_n be a suitable Lipschitz approximation of $\mathbf{1}_{A_n^{(q)}}$, defined as in 4.4. Then

$$\begin{split} \left| \int 1_{A_{n}^{(q)}} (1_{A_{n}^{(q)}} \circ T^{j}) \, dm - \left(\int 1_{A_{n}^{(q)}} \, dm \right)^{2} \right| & \leq \left| \int \Phi_{n} (\Phi_{n} \circ T^{j}) \, dm - \left(\int \Phi_{n} \, dm \right)^{2} \right| \\ & + \left| \left(\int \Phi_{n} \, dm \right)^{2} - \left(\int 1_{A_{n}^{(q)}} \, dm \right)^{2} \right| \\ & + \left| \int 1_{A_{n}^{(q)}} (1_{A_{n}^{(q)}} \circ T^{j}) \, dm - \int \Phi_{n} (\Phi_{n} \circ T^{j}) \, dm \right|. \end{split}$$

If $(\log n)^5 \leq j \leq \lfloor n/k_n \rfloor$, then, due to the decay of correlations, the first term is upper bounded by

$$\left| \int \Phi_n(\Phi_n \circ T^j) \, dm - \left(\int \Phi_n \, dm \right)^2 \right| \le C \, n^2 \, \theta^j \le C \, n^{-2}$$

if n is sufficiently large. Recalling Remark 2.1, for the second term we obtain, for n big enough,

$$\left| \left(\int \Phi_n \, dm \right)^2 - \left(\int 1_{A_n^{(q)}} \, dm \right)^2 \right| \le C \, m(A_n^{(q)} \setminus D_n) \le C \, n^{-2}.$$

Similarly, we estimate the third term as follows

$$\left| \int \Phi_n(\Phi_n \circ T^j) \, dm - \int 1_{A_n^{(q)}} (1_{A_n^{(q)}} \circ T^j) \, dm \right| \le 2 \, m(A_n^{(q)} \setminus D_n) \le C \, n^{-2}.$$

Hence equation (5.1) is satisfied.

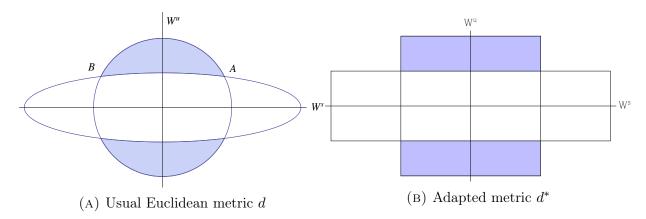


FIGURE 1. The sets $A_n^{(q)}$ for q > 0.

5.1. The periodic case: checking $\mathcal{A}_q'(u_n)$ for q>0. Suppose ζ is a periodic point of minimal period q for the map T. If $G=T^q$, then ζ is a fixed point for the Anosov diffeomorphism G, which has eigenvalues $\frac{1}{\lambda^q}$ and λ^q , with $\frac{1}{|\lambda|^q}<1<|\lambda|^q$, where $\frac{1}{\lambda}$ and λ are the eigenvalues of the original map T.

Note that $A_n^{(q)}$ corresponds to the shaded region in Figure 1a, the ball represents the set $\{X_0>u_n\}$, while the ellipse is the pre-image $T^{-q}(\{X_0>u_n\})$. To compute the diameter of $A_n^{(q)}$, we just need to find the distance between the points A and B on the intersection of the ball with the ellipse. If we use local coordinates x,y for the unstable and stable local manifolds through ζ , respectively, then the ball's equation is $x^2+y^2=s_n^2$ and the ellipse's equation is $|\lambda|^{2q}x^2+|\lambda|^{-2q}y^2=s_n^2$, where $s_n^2\sim \tau/(\pi n)$. Then the x coordinate of the points A and B is $\frac{s_n}{\sqrt{|\lambda|^{2q}+1}}$, while the y coordinates are $\frac{|\lambda|^q s_n}{\sqrt{|\lambda|^{2q}+1}}$ and its symmetric counterpart, respectively. This means that the diameter of $A_n^{(q)}$ is $\frac{2|\lambda|^q s_n}{\sqrt{|\lambda|^{2q}+1}}$.

Note that, under G^{-j} , the local stable manifolds in a small ball around ζ expand uniformly and the local unstable manifolds contract uniformly. By construction of the sets $A_n^{(q)}$ it follows that $G^{-j}(A_n^{(q)}) \cap A_n^{(q)} = \emptyset$, while the diameter of $G^{-j}(A_n^{(q)})$, which is measured along the stable direction, is less than 1. Hence, $G^{-j}(A_n^{(q)}) \cap A_n^{(q)} = \emptyset$ for all $j = 1, \ldots, g(n)$, where

$$g(n) = \left| \frac{\log n + \log(|\lambda|^{2q} + 1) - 2\log(2|\lambda|^{q}(\tau/\pi)^{1/2})}{2q \log |\lambda|} \right|.$$

Since ζ is periodic of period q, for large n we have $T^{-j}A_n^{(q)}\cap A_n^{(q)}=\emptyset$ for $j=1,\ldots,qg(n)$. Thus,

$$\sum_{j=1}^{qg(n)} m(A_n^{(q)} \cap T^{-j}(A_n^{(q)})) = 0.$$
 (5.2)

Let $x \in A_n^{(q)} \cap W_{\zeta}^u$ and $\gamma_x := W_x^s \cap A_n^{(q)}$. For j > qg(n), the curve $T^{-j}(\gamma_x)$ may be sufficiently long to wrap around the torus. Yet, if this is the case then

$$|T^{-j}(\gamma_x) \cap A_n^{(q)}| \le \frac{C}{\sqrt{n}} |T^{-j}(\gamma_x)|.$$

Indeed, the length of $A_n^{(q)}$ in the stable direction is bounded above by $C_1 \frac{1}{\sqrt{n}}$, for some constant C_1 independent of the local stable manifold. As T restricted to local stable manifolds has bounded distortion (in fact its derivative at any point is even constant, equal to L), it follows that $|\{y: y \in \gamma_x, T^{-j}(y) \in A_n^{(q)}\}| \leq \frac{C}{\sqrt{n}} |\gamma_x|$. Moreover, since the 2-dimensional Lebesgue measure m is invariant by T, we have

$$m\left(A_{n}^{(q)} \cap T^{-j}(A_{n}^{(q)})\right) = m\left(A_{n}^{(q)} \cap T^{j}(A_{n}^{(q)})\right) = \int_{A_{n}^{(q)} \cap W_{\zeta}^{u}} |\gamma_{x} \cap T^{j}(A_{n}^{(q)})| dm_{x}$$

$$\leq \frac{C}{\sqrt{n}} \int_{A_{n}^{(q)} \cap W_{\zeta}^{u}} |\gamma_{x}| dm_{x} \leq \frac{C}{\sqrt{n}} m(A_{n}^{(q)}).$$

This implies that

$$n \sum_{j=qg(n)+1}^{\log^5 n} m\left(A_n^{(q)} \cap T^{-j}(A_n^{(q)})\right) = O\left(n(\log n)^5 n^{-1/2} m(A_n^{(q)})\right) = O(n^{-1/2} \log^5 n) \xrightarrow[n \to \infty]{} 0$$
(5.3)

which, together with (5.1) and (5.2), yields (3.4).

5.2. The non-periodic case: checking $\mathcal{I}'_q(u_n)$ for q=0. Let us consider now a non-periodic point ζ . Note that, in this case, $A_n^{(q)}=U_n=\{X_0>u_n\}$ is a ball centered at ζ such that $m(U_n)\sim \tau/n$. We split the sum in (3.4) into three parts:

$$n \sum_{j=0}^{\lfloor n/k_n \rfloor} m(U_n \cap T^{-j}(U_n)) = n \sum_{j=0}^{g(n)} m(U_n \cap T^{-j}(U_n)) + n \sum_{j=g(n)+1}^{\log^5 n} m(U_n \cap T^{-j}(U_n)) + n \sum_{j=\log^5 n+1}^{n/k_n} m(U_n \cap T^{-j}(U_n)),$$

$$(5.4)$$

where $g(n) = \lfloor \frac{\log n - \log \pi}{2 \log |\lambda|} \rfloor$ is such that, for j > g(n), the set $T^{-j}(U_n)$ is sufficiently stretched along the stable direction to start wrapping around the torus. The third summand on the right of (5.4) goes to 0 on account of (5.1). We proceed considering the other summands, starting by the first. Contrary to the periodic setting, we now have to deal with the possibility of $U_n \cap T^{-j}(U_n) \neq \emptyset$ for $j \leq g(n)$. Let $R_n = \min\{j \in \mathbb{N} : U_n \cap T^{-j}(U_n) \neq \emptyset\}$. By continuity of T and the fact that $\bigcap_n U_n = \{\zeta\}$, we have that $R_n \to \infty$ as $n \to \infty$.

For $j \leq g(n)$, observe that $T^{-j}(U_n)$ is an ellipse centered at $T^{-j}(\zeta)$, stretched along the stable direction by a factor $|\lambda|^j$ and contracted on the unstable direction by $1/|\lambda|^j$. Also,

notice that, as we are dealing with a projection on the torus of a linear map in \mathbb{R}^2 , the stable and unstable directions are always lined up regardless of the position of $T^{-j}(\zeta)$. Now, since the set $T^{-j}(U_n)$ is not long enough (on the stable direction) to wrap around the torus, the diameter of U_n is $2s_n$, where $s_n \sim \sqrt{\tau}/\sqrt{\pi n}$. The length of $T^{-j}(U_n)$ along the stable direction (which coincides with the diameter of $T^{-j}(U_n)$) is $|\lambda|^j s_n$; and the length of $T^{-j}(U_n)$ on the unstable direction is $|\lambda|^{-j} s_n$. So, by the geometry of the sets, the area of the intersection $U_n \cap T^{-j}(U_n)$ occupies a portion of $T^{-j}(U_n)$ corresponding at most to a constant times $|\lambda|^{-j}$ of its area. That is, there exists some C > 0 independent of j such that

$$m(U_n \cap T^{-j}(U_n)) \le C |\lambda|^{-j} m(T^{-j}(U_n)) = C |\lambda|^{-j} m(U_n).$$

In the case of the Euclidean metric we can take $C \approx 2/\pi$. Then it follows that

$$n \sum_{j=0}^{g(n)} m(U_n \cap T^{-j}(U_n)) = n \sum_{j=R_n}^{g(n)} |\lambda|^{-j} \tau / n \le C \tau |\lambda|^{-R_n} \xrightarrow[n \to \infty]{} 0.$$

We are left to estimate the second summand of (5.4). Let $x \in U_n \cap W_{\zeta}^u$ and $\gamma_x := W_x^s \cap U_n$. For j > g(n), the curve $T^{-j}(\gamma_x)$ may be sufficiently long to wrap around the torus. Set

$$\tilde{U}_n^j := \bigcup_{\left\{x \in U_n \cap W_\zeta^u : |T^{-j}(\gamma_x)| > 1/2\right\}} \gamma_x.$$

Note that for the usual Euclidean metric we have $|W_{\zeta}^{u} \cap U_{n} \setminus \tilde{U}_{n}^{j}| \leq 2\sqrt{s_{n}^{2} - |\lambda|^{-2j}/16}$. Since for $x \in \tilde{U}_{n}^{j}$, we have $|T^{-j}(\gamma_{x})| > 1/2$ and the diameter of U_{n} is less than $2s_{n}$, then there exists C > 0 such that, for all j > g(n) and all $x \in \tilde{U}_{n}^{j}$, we have

$$|T^{-j}(\gamma_x) \cap U_n| \le C s_n |T^{-j}(\gamma_x)|.$$

As T restricted to local stable manifolds has bounded distortion (since the derivative is constant), it follows that $|\{y: y \in \gamma_x, T^{-j}(y) \in U_n\}| \leq C s_n |\gamma_x|$. Moreover, by the T-invariance of m we have

$$m(U_n \cap T^{-j}(\tilde{U}_n^j)) = m(\tilde{U}_n^j \cap T^j(U_n)) = \int_{\tilde{U}_n^j \cap W_\zeta^u} |\gamma_x \cap T^j(A_n^{(q)})| \, dm_x$$

$$\leq C \, s_n \int_{\tilde{U}_n^j \cap W_\zeta^u} |\gamma_x| \, dm_x \leq C \, s_n \, m(\tilde{U}_n^j). \tag{5.5}$$

Now, we focus on $T^{-j}(U_n \setminus \tilde{U}_n^j)$ which is a made of two connected components with diameter 1/2 along the stable direction and with length $|\lambda|^{-j} s_n \left(1 - \sqrt{1 - |\lambda|^{-2j}/(16s_n^2)}\right)$. As before, the geometry dictates that the area of the intersection $U_n \cap T^{-j}(U_n \setminus \tilde{U}_n^j)$ occupies a portion of $T^{-j}(U_n \setminus \tilde{U}_n^j)$ that corresponds at most to a constant times s_n of its area. Indeed, there exists C > 0 such that,

$$m(U_n \cap T^{-j}(U_n \setminus \tilde{U}_n^j)) \le C \, s_n \, m(T^{-j}(U_n \setminus \tilde{U}_n^j)) = C \, s_n \, m(U_n \setminus \tilde{U}_n^j). \tag{5.6}$$

Joining (5.5), (5.6) and recalling that $s_n \sim \sqrt{\tau}/\sqrt{\pi n}$, we deduce that

$$n\sum_{j=q(n)+1}^{\log^5 n} m(U_n \cap T^{-j}(U_n)) = O(n(\log n)^5 n^{-1/2} m(U_n)) = O(n^{-1/2} \log^5 n) \xrightarrow[n \to \infty]{} 0. \quad (5.7)$$

6. Computing the Extremal Index

We recall that the Extremal Index ϑ is given by the limit in (3.5).

In the case of a non-periodic point ζ , we have q=0 and the computation is trivial because $A_n^{(q)}=U_n$, hence $\vartheta=1$.

If ζ is a periodic point of prime period q, we need to compute the area of the shaded regions in Figure 1a. Recalling that the second coordinate of the point A of the picture is $y_A = \frac{|\lambda|^q s_n}{\sqrt{|\lambda|^2 q_{+1}}}$ and using the symmetry of the region, we have

$$m(A_n^{(q)}) = 4 \int_0^{y_A} \sqrt{s_n^2 - y^2} - |\lambda|^{-q} \sqrt{s_n^2 - y^2} |\lambda|^{-2q} dy$$

$$= 2 s_n^2 \left(\arcsin \frac{|\lambda|^q}{\sqrt{|\lambda|^{2q} + 1}} - \arcsin \frac{1}{\sqrt{|\lambda|^{2q} + 1}} \right).$$
(6.1)

Thus, in this case,

$$\vartheta_q = \frac{2}{\pi} \left(\arcsin \frac{|\lambda|^q}{\sqrt{|\lambda|^{2q} + 1}} - \arcsin \frac{1}{\sqrt{|\lambda|^{2q} + 1}} \right). \tag{6.2}$$

Remark 6.1. It is interesting to observe that the metric used affects the value of the Extremal Index. In fact, if we had used the adapted metric, so that its balls would correspond to squares with sides lined up with the stable and unstable directions, then the EI would be equal to $1 - |\lambda|^{-q}$, which is unsurprisingly more in tune with the one-dimensional setting. See Figure 1b. Notice also that, in both metrics, the extremal index approaches the non-periodic value as the period goes to ∞ , that is, $\lim_{q\to +\infty} \vartheta_q = 1$.

7. Conclusion

For C^2 expanding maps of the interval a dichotomy exists between the extremal behavior of periodic points and non-periodic points [7, 11, 22]. In this paper we demonstrated that the same dichotomy holds for hyperbolic total automorphisms. The extreme value statistics of observations maximized at generic points in a variety of non-uniformly hyperbolic systems is known to be the same as that of a sequence of iid random variables with the same distribution function [9, 10, 19, 16] but the statistics of functions maximized at periodic points has not yet been established. We believe that it will be the same as in the iid case for smooth non-uniformly hyperbolic systems but expect mixed distributions (as in [1]) for non-uniformly systems with discontinuities.

8. RARE EVENTS POINT PROCESSES

In order to make the presentation more tractable we started by stating the dichotomy in terms of the possible limit distributions for the maxima when the centers are chosen to be either non-periodic or periodic points. However, we can enrich the results without much extra work by considering the convergence of point processes of exceedances or rare events.

If we consider multiple exceedances we are led to point processes of rare events counting the number of exceedances in a certain time frame. For every $A \subset \mathbb{R}$ we define

$$\mathscr{N}_u(A) := \sum_{i \in A \cap \mathbb{N}_0} \mathbf{1}_{X_i > u}.$$

In the particular case where A = [a, b) we simply write $\mathcal{N}_{u,a}^b := \mathcal{N}_u([a, b))$. Observe that $\mathcal{N}_{u,0}^n$ counts the number of exceedances amongst the first n observations of the process X_0, X_1, \ldots, X_n or, in other words, the number of entrances in U(u) up to time n. Also, note that

$$\{\mathcal{N}_{u,0}^n = 0\} = \{M_n \le u\} \tag{8.1}$$

In order to define a point process that captures the essence of an EVL and HTS through (8.1), we need to re-scale time using the factor $v := 1/\mathbb{P}(X > u)$ given by Kac's Theorem. Let S denote the semi-ring of subsets of \mathbb{R}_0^+ whose elements are intervals of the type [a, b), for $a, b \in \mathbb{R}_0^+$. Denote by \mathcal{R} the ring generated by S; recall that for every $J \in \mathcal{R}$ there are $k \in \mathbb{N}$ and k intervals $I_1, \ldots, I_k \in S$, say $I_j = [a_j, b_j)$ with $a_j, b_j \in \mathbb{R}_0^+$, such that $J = \bigcup_{i=1}^k I_j$. For $I = [a, b) \in S$ and $\alpha \in \mathbb{R}$, we set $\alpha I := [\alpha a, \alpha b)$ and $I + \alpha := [a + \alpha, b + \alpha)$. Similarly, for $J \in \mathcal{R}$, we define $\alpha J := \alpha I_1 \cup \cdots \cup \alpha I_k$ and $J + \alpha := (I_1 + \alpha) \cup \cdots \cup (I_k + \alpha)$.

Definition 1. Given $J \in \mathcal{R}$ and sequences $(u_n)_{n \in \mathbb{N}}$ and $(v_n)_{n \in \mathbb{N}}$, the rare event point process (REPP) is defined by counting the number of exceedances (or hits to $U(u_n)$) during the (re-scaled) time period $v_n J \in \mathcal{R}$. More precisely, for every $J \in \mathcal{R}$, set

$$N_n(J) := \mathcal{N}_{u_n}(v_n J) = \sum_{j \in v_n J \cap \mathbb{N}_0} \mathbf{1}_{X_j > u_n}. \tag{8.2}$$

Under dependence conditions similar to the ones previously analyzed, the REPP just defined converges in distribution to a standard Poisson process when no clustering is involved, and converges in distribution to a compound Poisson process with intensity ϑ and a geometric multiplicity distribution function otherwise. For the sake of completeness, we now define what we mean by a Poisson and a compound Poisson process. (See [20] for more details.)

Definition 2. Let $Y_1, Y_2, ...$ be an iid sequence of random variables with common exponential distribution of mean $1/\vartheta$. Let $Z_1, Z_2, ...$ be another iid sequence of random variables, independent of the previous one, and with distribution function π . Given these sequences, for $J \in \mathcal{R}$, set

$$N(J) = \int \mathbf{1}_J d\left(\sum_{i=1}^{\infty} Z_i \delta_{Y_1 + \dots + Y_i}\right),\,$$

where δ_t denotes the Dirac measure at t > 0. Whenever we are in this setting, we say that N is a compound Poisson process of intensity ϑ and multiplicity d.f. π .

Remark 8.1. In this paper, the multiplicity will always be integer valued. This means that π is completely defined by the values $\pi_k = \mathbb{P}(Z_1 = k)$, for every $k \in \mathbb{N}_0$. Note that, if $\pi_1 = 1$ and $\vartheta = 1$, then N is the standard Poisson process and, for every t > 0, the random variable N([0,t)) has a Poisson distribution of mean t.

Remark 8.2. When clustering is involved, typically π is a geometric distribution of parameter $\vartheta \in (0,1]$, i.e., $\pi_k = \vartheta(1-\vartheta)^{k-1}$, for every $k \in \mathbb{N}_0$. This means that, as in [17], the random variable N([0,t)) follows a Pólya-Aeppli distribution

$$\mathbb{P}(N([0,t))) = k = e^{-\vartheta t} \sum_{j=1}^{k} \vartheta^{j} (1-\vartheta)^{k-j} \frac{(\vartheta t)^{j}}{j!} {k-1 \choose j-1},$$

for all $k \in \mathbb{N}$ and $\mathbb{P}(N([0,t))) = 0 = e^{-\vartheta t}$.

We are now in condition to state the dichotomy in terms of the convergence of REPP.

Theorem 8.3. Let $T: \mathbb{T}^2 \to \mathbb{T}^2$ be an Anosov linear diffeomorphism. Fix $\zeta \in \mathbb{T}^2$ and define the stochastic process X_0, X_1, X_2, \ldots as in (2.5). Let $(u_n)_{n \in \mathbb{N}}$ be a sequence satisfying (2.6) for some $\tau \geq 0$ and $(v_n)_{n \in \mathbb{N}}$ be given by $v_n = 1/\mathbb{P}(X_0 > u_n)$. Consider the REPP N_n as in Definition 2. Then

- (1) if ζ is not periodic, the REPP N_n converges in distribution to the standard Poisson process.
- (2) if ζ is periodic of prime period q, the REPP N_n converges in distribution to a compound Poisson process with intensity ϑ given by (6.2) and multiplicity d.f. π given by (8.6).

As in the computation of the EI, the metric used has an important impact on the multiplicity distribution obtained when we have convergence of the REPP to a compound Poisson process. Indeed, a similar result for the adapted metric d^* given in (2.7) is as follows.

Theorem 8.4. Let $T: \mathbb{T}^2 \to \mathbb{T}^2$ be an Anosov linear diffeomorphism. Fix $\zeta \in \mathbb{T}^2$ and define the stochastic process X_0, X_1, X_2, \ldots as in (2.5) for the observable φ in (2.4) but evaluated with the metric d^* instead. Let $(u_n)_{n \in \mathbb{N}}$ be a sequence satisfying (2.6) for some $\tau \geq 0$ and $(v_n)_{n \in \mathbb{N}}$ be given by $v_n = 1/\mathbb{P}(X_0 > u_n)$. Consider the REPP N_n as in Definition 1. Then

- (1) if ζ is not periodic, the REPP N_n converges in distribution to the standard Poisson process.
- (2) if ζ is periodic of prime period q, the REPP N_n converges in distribution to a compound Poisson process with intensity $\vartheta = 1 |\lambda|^{-q}$ and geometric multiplicity d.f. π^* given by $\pi^*(\kappa) = \vartheta(1 \vartheta)^{\kappa 1}$, for all $\kappa \in \mathbb{N}$.

8.1. **Absence of clustering.** When condition $\mathcal{A}'_0(u_n)$ holds, there is no clustering and so we may benefit from a criterion, proposed by Kallenberg [20, Theorem 4.7], which applies only to simple point processes without multiple events. Accordingly, we can merely adjust condition $\mathcal{A}_0(u_n)$ to this scenario of multiple exceedances in order to prove that the REPP converges in distribution to a standard Poisson process. We denote this adapted condition by:

Condition $(D_3(u_n))$. Let $A \in \mathcal{R}$ and $t \in \mathbb{N}$. We say that $D_3(u_n)$ holds for the sequence X_0, X_1, \ldots if

$$|\mathbb{P}(\{X_0 > u_n\} \cap \{\mathcal{N}_{u_n}(A+t) = 0\}) - \mathbb{P}(\{X_0 > u_n\})\mathbb{P}(\mathcal{N}_{u_n}(A) = 0)| \le \gamma(n,t),$$

where $\gamma(n,t)$ is nonincreasing in t for each n and $n\gamma(n,t_n) \to 0$ as $n \to \infty$ for some sequence $t_n = o(n)$. (The last equality means that $t_n/n \to 0$ as $n \to \infty$).

In [9, Theorem 5] it was proved a strengthening of [8, Theorem 1] which essentially says that, under conditions $D_3(u_n)$ and $\mathcal{J}'_0(u_n)$, the REPP N_n defined in (8.2) converges in distribution to a standard Poisson process.

The proof of condition $D_3(u_n)$ follows after minor adjustments to the proof of $\mathcal{L}_q(u_n)$ in Section 4. Since condition $\mathcal{L}'_0(u_n)$ holds at every non-periodic point ζ (see Section 5.2), then for all such points ζ the corresponding REPP N_n converges in distribution to a standard Poisson process.

8.2. **Presence of clustering.** Condition $\mathcal{I}'_0(u_n)$ prevents the existence of clusters of exceedances, which implies that the EVL is standard exponential $\bar{H}(\tau) = e^{-\tau}$. When $\mathcal{I}'_0(u_n)$ fails, the clustering of exceedances is responsible for the appearance of a parameter $0 < \vartheta < 1$ in the EVL, called EI, and implies that, in this case, $\bar{H}(\tau) = e^{-\vartheta \tau}$. In [10], the authors established a connection between the existence of an EI less than 1 and a periodic behavior. This was later generalized for REPP in [11].

For the convergence of the REPP when there is clustering, one cannot use the aforementioned criterion of Kallenberg because the point processes are not simple anymore and possess multiple events. This means that a much deeper analysis must be done in order to obtain convergence of the REPP. This was carried out in [11] and we describe below the main results and conditions needed.

Let ζ be a periodic point of prime period q. Firstly, we consider the sequence $(U^{(\kappa)}(u_n))_{\kappa \geq 0}$ of nested balls centred at ζ given by

$$U^{(0)}(u_n) = U(u_n) = U_n$$
 and $U^{(\kappa)}(u_n) = T^{-q}(U^{(\kappa-1)}(u_n)) \cap U(u_n)$, for all $\kappa \in \mathbb{N}$. (8.3)

Then, for $i, \kappa, \ell, s \in \mathbb{N} \cup \{0\}$, we define the following events:

$$Q_{q,i}^{\kappa}(u_n) := T^{-i} \left(U^{(\kappa)}(u_n) - U^{(\kappa+1)}(u_n) \right). \tag{8.4}$$

Note that $Q_{q,0}^0(u_n) = A_n^{(q)}$. Besides, $U_n = \bigcup_{\kappa=0}^{\infty} Q_{q,0}^{\kappa}(u_n)$, which means that the ball centred at ζ which corresponds to U_n can be decomposed into a sequence of disjoint strips where

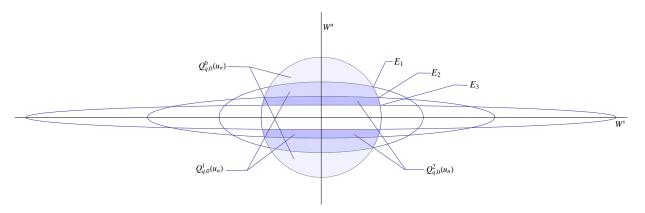


FIGURE 2. The sets $Q_{q,0}^{\kappa}(u_n)$ for the Euclidean metric.

 $Q_{q,0}^0(u_n)$ are the most outward strips and the inner strips $Q_{q,0}^{\kappa+1}(u_n)$ are sent outward by T^p onto the strips $Q_{q,0}^{\kappa}(u_n)$, i.e., $T^q(Q_{q,0}^{\kappa+1}(u_n)) = Q_{q,0}^{\kappa}(u_n)$.

We are now ready to state the adapted condition:

Condition $(D_q(u_n)^*)$. We say that $D_q(u_n)^*$ holds for the sequence X_0, X_1, X_2, \ldots if for any integers $t, \kappa_1, \ldots, \kappa_{\zeta}$, n and any $J = \bigcup_{i=2}^q I_j \in \mathcal{R}$ with $\inf\{x : x \in J\} \geq t$,

$$\left| \mathbb{P}\left(Q_{q,0}^{\kappa_1}(u_n) \cap \left(\cap_{j=2}^q \mathcal{N}_{u_n}(I_j) = \kappa_j \right) \right) - \mathbb{P}\left(Q_{q,0}^{\kappa_1}(u_n) \right) \mathbb{P}\left(\cap_{j=2}^q \mathcal{N}_{u_n}(I_j) = \kappa_j \right) \right| \leq \gamma(n,t),$$

where for each n we have that $\gamma(n,t)$ is nonincreasing in t and $n\gamma(n,t_n) \to 0$ as $n \to \infty$, for some sequence $t_n = o(n)$.

In [11], for technical reasons only, the authors introduced a slight modification to $\mathcal{L}'_q(u_n)$. This condition was denoted by $D'_q(u_n)^*$ and it requires that

$$\lim_{n \to \infty} n \sum_{j=1}^{[n/k_n]} \mathbb{P}(Q_{q,0}(u_n) \cap \{X_j > u_n\}) = 0,$$

which holds whenever condition $\coprod_{q}'(u_n)$ does.

From the study developed in [11], it follows that if X_0, X_1, \ldots satisfies conditions $D_q(u_n)^*$, $D'_q(u_n)^*$ and $\lim_{n\to\infty} \sum_{\kappa\geq 1} \mathbb{P}(U^{(\kappa)}(u_n)) = 0$, where $(u_n)_{n\in\mathbb{N}}$ is such that (2.6) holds, then the REPP N_n converges in distribution to a compound Poisson process with intensity ϑ given by (3.5) and multiplicity d.f. π given by:

$$\pi(\kappa) = \lim_{n \to \infty} \frac{\left(\mathbb{P}(Q_{q,0}^{\kappa - 1}(u_n)) - \mathbb{P}(Q_{q,0}^{\kappa}(u_n)) \right)}{\mathbb{P}(Q_{q,0}^0(u_n))}. \tag{8.5}$$

8.3. Computing the multiplicity distribution. In order to compute the multiplicity distribution, we need to calculate $m(Q_{q,0}^{\kappa})$. To do so, let E_i denote one of the four intersection points of $\partial T^{-iq}(U_n)$ with ∂U_n . Let (x_{E_i}, y_{E_i}) denote the coordinates of E_i on the unstable and stable directions, which correspond to the vertical and horizontal directions,

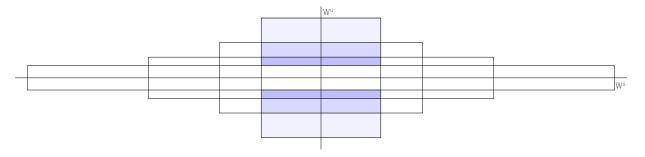


FIGURE 3. The sets $Q_{q,0}^{\kappa}(u_n)$ in the adapted metric.

respectively, in Figure 2. We assume that the choice of the E_i 's among the possible four points is made so that $x_{E_i}, y_{E_i} > 0$. Hence,

$$y_{E_i} = \frac{s_n |\lambda|^{iq}}{\sqrt{1 + \lambda^{2iq}}},$$

because $\partial T^{-iq}(U_n)$ is described by the equation $\lambda^{2iq}x^2 + \lambda^{-2iq}y^2 = s_n^2$. Therefore, for $\kappa \geq 1$, we may write

$$m(Q_{q,0}^{\kappa}(u_n)) = 4 \int_0^{y_{E_i}} \lambda^{-2iq} \sqrt{s_n^2 - \lambda^{-2iq} y^2} dy + 4 \int_{y_{E_i}}^{y_{E_{i+1}}} \sqrt{s_n^2 - y^2} dy$$

$$-4 \int_0^{y_{E_{i+1}}} \lambda^{-2(i+1)q} \sqrt{s_n^2 - \lambda^{-2(i+1)q} y^2} dy$$

$$= 2s_n^2 \left(\arcsin \frac{|\lambda|^{(i+1)q}}{\sqrt{1 + \lambda^{2(i+1)q}}} - \arcsin \frac{1}{\sqrt{1 + \lambda^{2(i+1)q}}} \right)$$

$$+ 2s_n^2 \left(\arcsin \frac{1}{\sqrt{1 + \lambda^{2iq}}} - \arcsin \frac{|\lambda|^{iq}}{\sqrt{1 + \lambda^{2iq}}} \right)$$

Recall that $Q_{q,0}^0(u_n) = A_n^{(q)}$ and its measure has been computed in (6.1). Hence, using formula (8.5), we obtain

$$\pi(\kappa) = \frac{\left(\arcsin\frac{1}{\sqrt{1+\lambda^{2(\kappa-1)q}}} - \arcsin\frac{|\lambda|^{(\kappa-1)q}}{\sqrt{1+\lambda^{2(\kappa-1)q}}}\right) + \left(\arcsin\frac{1}{\sqrt{1+\lambda^{2(\kappa+1)q}}} - \arcsin\frac{|\lambda|^{(\kappa+1)q}}{\sqrt{1+\lambda^{2(\kappa+1)q}}}\right)}{\arcsin\frac{|\lambda|^q}{\sqrt{1+\lambda^{2q}}} - \arcsin\frac{1}{\sqrt{1+\lambda^{2q}}}} + \frac{2\left(\arcsin\frac{|\lambda|^{\kappa q}}{\sqrt{1+\lambda^{2\kappa q}}} - \arcsin\frac{1}{\sqrt{1+\lambda^{2\kappa q}}}\right)}{\arcsin\frac{|\lambda|^q}{\sqrt{1+\lambda^{2q}}} - \arcsin\frac{1}{\sqrt{1+\lambda^{2q}}}}$$

$$(8.6)$$

Remark 8.5. As for the EI (see Remark 6.1), the multiplicity distribution depends on the metric used. If we were to use the adapted metric d^* , in which balls correspond to squares whose sides are lined up with the stable and unstable directions, then we would obtain a geometric distribution for the multiplicity, i.e., $\pi^*(\kappa) = \vartheta^*(1-\vartheta^*)^{\kappa-1}$, where $\vartheta^* = 1-|\lambda|^{-q}$.

Observe also that in the Euclidean metric case we do not obtain a geometric distribution, as one did for repelling periodic points (see [11]) or in the case of the adapted metric mentioned in Remark 8.5. However, taking into account that

$$\arcsin\left(\frac{x}{\sqrt{1+x^2}}\right) = \frac{\pi}{2} - \frac{1}{x} + o(1/x) \quad \text{and} \quad \arcsin\left(\frac{1}{\sqrt{1+x^2}}\right) = \frac{1}{x} + o(1/x),$$

one deduces that

$$\lim_{\kappa \to \infty} \frac{\pi(\kappa+1)}{\pi(\kappa)} = |\lambda|^{-q} = \frac{\pi^*(\kappa+1)}{\pi^*(\kappa)},$$

where $\pi^*(\kappa)$ is as given in Theorem 8.4. Hence, we obtained a multiplicity distribution different from the geometric distribution but which is not that far from it.

The proof of condition $D_q(u_n)^*$ follows with some adjustments from the proof of $\mathcal{A}_q(u_n)$ in Section 4. Since condition $\mathcal{A}'_q(u_n)$ holds at every periodic point ζ , as was shown in Section 5.2, then condition $D'_q(U_n)^*$ from [11] also holds for all such points ζ . Let C_n^{κ} denote a rectangle, centered at ζ , with its sides lined up with the stable and unstable direction, of length s_n on the stable direction and $|\lambda|^{-\kappa q}s_n$ on the unstable direction. Observe that $U^{(\kappa)}(u_n) \subset C_n^{\kappa}$, which implies that

$$m(U^{(\kappa)}(u_n)) \le |\lambda|^{-\kappa q} s_n^2$$

and consequently

$$\lim_{n \to \infty} \sum_{\kappa > 1} m(U^{(\kappa)}(u_n)) = \lim_{n \to \infty} \frac{1}{1 - |\lambda|^{-q}} s_n^2 = 0.$$

So, at every periodic point ζ of prime period q, the respective REPP N_n converges in distribution to a compound Poisson process with intensity ϑ given by (6.2) and multiplicity d.f. π given by (8.6).

REFERENCES

- [1] H. Aytaç, J. M. Freitas, and S. Vaienti. Laws of rare events for deterministic and random dynamical systems. *Trans. Amer. Math.*, Published online with DOI: 10.1090/S0002-9947-2014-06300-9, 2014.
- [2] R. Bowen. Equilibrium states and the ergodic theory of Anosov diffeomorphisms. Lecture Notes in Mathematics, Vol. 470. Springer-Verlag, Berlin, 1975.
- [3] J.-R. Chazottes and P. Collet. Poisson approximation for the number of visits to balls in non-uniformly hyperbolic dynamical systems. *Ergodic Theory Dynam. Systems*, 33(1):49–80, 2013.
- [4] P. Collet. Statistics of closest return for some non-uniformly hyperbolic systems. *Ergodic Theory Dynam. Systems*, 21(2):401–420, 2001.
- [5] M. Denker, M. Gordin, and A. Sharova. A Poisson limit theorem for toral automorphisms. *Illinois J. Math.*, 48(1):1–20, 2004.
- [6] D. Dolgopyat. Limit theorems for partially hyperbolic systems. *Trans. Amer. Math. Soc.*, 356(4):1637–1689 (electronic), 2004.
- [7] A. Ferguson and M. Pollicott. Escape rates for gibbs measures. *Ergod. Theory Dynam. Systems*, 32:961–988, 2012.
- [8] A. C. M. Freitas and J. M. Freitas. On the link between dependence and independence in extreme value theory for dynamical systems. *Statist. Probab. Lett.*, 78(9):1088–1093, 2008.

- [9] A. C. M. Freitas, J. M. Freitas, and M. Todd. Hitting time statistics and extreme value theory. *Probab. Theory Related Fields*, 147(3-4):675–710, 2010.
- [10] A. C. M. Freitas, J. M. Freitas, and M. Todd. The extremal index, hitting time statistics and periodicity. *Adv. Math.*, 231(5):2626–2665, 2012.
- [11] A. C. M. Freitas, J. M. Freitas, and M. Todd. The compound Poisson limit ruling periodic extreme behaviour of non-uniformly hyperbolic dynamics. *Comm. Math. Phys.*, 321(2):483–527, 2013.
- [12] A. C. M. Freitas, J. M. Freitas, and M. Todd. Speed of convergence for laws of rare events and escape rates. *Stochastic Process. Appl.*, Published online with DOI: 10.1016/j.spa.2014.11.011, 2014.
- [13] J. M. Freitas, N. Haydn, and M. Nicol. Convergence of rare event point processes to the Poisson process for planar billiards. *Nonlinearity*, 27(7):1669–1687, 2014.
- [14] J. Galambos. The asymptotic theory of extreme order statistics. John Wiley & Sons, New York-Chichester-Brisbane, 1978. Wiley Series in Probability and Mathematical Statistics.
- [15] C. Gupta. Extreme-value distributions for some classes of non-uniformly partially hyperbolic dynamical systems. *Ergodic Theory Dynam. Systems*, 30(3):757–771, 2010.
- [16] C. Gupta, M. Holland, and M. Nicol. Extreme value theory and return time statistics for dispersing billiard maps and flows, Lozi maps and Lorenz-like maps. Ergodic Theory Dynam. Systems, 31(5):1363–1390, 2011.
- [17] N. Haydn and S. Vaienti. The compound Poisson distribution and return times in dynamical systems. *Probab. Theory Related Fields*, 144(3-4):517–542, 2009.
- [18] M. Hirata. Poisson law for Axiom A diffeomorphisms. Ergodic Theory Dynam. Systems, 13(3):533-556, 1993.
- [19] M. Holland, M. Nicol, and A. Török. Extreme value theory for non-uniformly expanding dynamical systems. *Trans. Amer. Math. Soc.*, 364(2):661–688, 2012.
- [20] O. Kallenberg. Random measures. Akademie-Verlag, Berlin, fourth edition, 1986.
- [21] A. Katok and B. Hasselblatt. *Introduction to the modern theory of dynamical systems*, volume 54 of *Encyclopedia of Mathematics and its Applications*. Cambridge University Press, Cambridge, 1995. With a supplementary chapter by Katok and Leonardo Mendoza.
- [22] G. Keller. Rare events, exponential hitting times and extremal indices via spectral perturbation. *Dyn. Syst.*, 27(1):11–27, 2012.
- [23] M. R. Leadbetter. On extreme values in stationary sequences. Z. Wahrscheinlichkeitstheorie und Verw. Gebiete, 28:289–303, 1973/74.
- [24] M. R. Leadbetter, G. Lindgren, and H. Rootzén. Extremes and related properties of random sequences and processes. Springer Series in Statistics. Springer-Verlag, New York, 1983.
- [25] A. Manning. There are no new Anosov diffeomorphisms on tori. Amer. J. Math., 96:422–429, 1974.
- [26] S. I. Resnick. Extreme values, regular variation, and point processes, volume 4 of Applied Probability. A Series of the Applied Probability Trust. Springer-Verlag, New York, 1987.
- [27] L. Zhang. Borel cantelli lemmas and extreme value theory for geometric Lorenz models. Preprint arXiv:1404.1656, 2014.

Maria Carvalho, Centro de Matemática & Faculdade de Ciências da Universidade do Porto, Rua do Campo Alegre 687, 4169-007 Porto, Portugal

E-mail address: mpcarval@fc.up.pt

Ana Cristina Moreira Freitas, Centro de Matemática & Faculdade de Economia da Universidade do Porto, Rua Dr. Roberto Frias, 4200-464 Porto, Portugal

E-mail address: amoreira@fep.up.pt

 URL : http://www.fep.up.pt/docentes/amoreira/

JORGE MILHAZES FREITAS, CENTRO DE MATEMÁTICA & FACULDADE DE CIÊNCIAS DA UNIVERSIDADE DO PORTO, RUA DO CAMPO ALEGRE 687, 4169-007 PORTO, PORTUGAL

E-mail address: jmfreita@fc.up.pt

URL: http://www.fc.up.pt/pessoas/jmfreita/

MARK HOLLAND, DEPARTMENT OF MATHEMATICS (CEMPS), HARRISON BUILDING (327), NORTH PARK ROAD, EXETER, EX4 4QF, UK

 $E ext{-}mail\ address: M.P.Holland@exeter.ac.uk}$

URL: http://empslocal.ex.ac.uk/people/staff/mph204/

MATTHEW NICOL, DEPARTMENT OF MATHEMATICS, UNIVERSITY OF HOUSTON, HOUSTON, TX 77204, USA

E-mail address: nicol@math.uh.edu

 URL : http://www.math.uh.edu/~nicol/